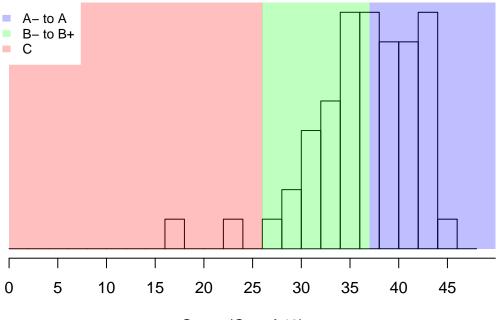
# Exam 2 3/26/19

There are five questions, each of which has several parts. Neither the questions nor the parts are necessarily in order from easiest to most difficult. Make sure you have taken a look at and attempted all of the questions in the allotted time. Stop working and immediately turn in your exam when time has been called.

Name:		
Question	Maximum Possible Points	
1	14	
2	9	
3	6	
4	8	
5	10	
Total	47	

The figure below shows a histogram of scores. The mean was 36.78, the median was 37.5, and the standard deviation was 5.53. The rank correlation between the first and second exam scores was 0.58.

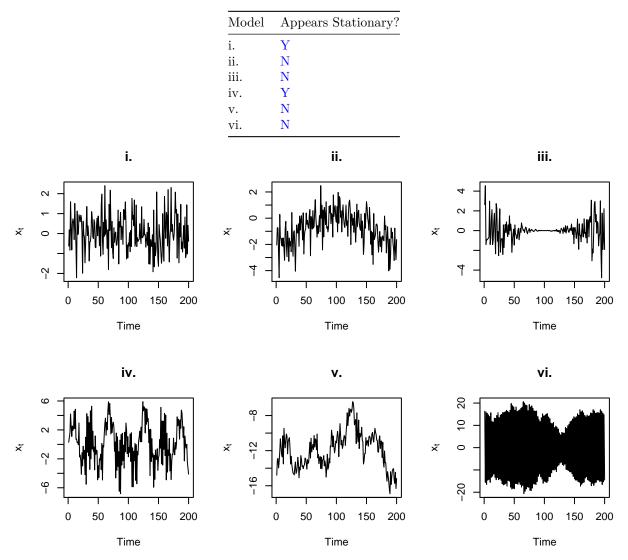


# Exam 2

Score (Out of 49)

### 1. Assessing Stationarity

(a) Based on the following time series plots, indicate whether or not each time series appears to be stationary by filling in "yes" or "no."



(b) For each of the following models, indicate whether or not the model is approximately consistent with the null hypothesis of an augmented Dickey-Fuller (ADF) test, either with or without a linear time trend, by filling in "yes" or "no." Recall that we can write we can write any  $\mathbf{ARMA}(p,q)$  as approximately an  $\mathbf{AR}(k)$  process for some value of k. Then the approximate null hypothesis of an augmented ADF test without trend is that  $x_t - x_{t-1}$  is a stationary  $\mathbf{ARMA}(p,q)$  process, whereas the null hypothesis of an augmented ADF test with trend is  $x_t - x_{t-1} - bt$  is a stationary  $\mathbf{ARMA}(p,q)$  process for constant b.

- i.  $x_t = w_t, w_t \stackrel{i.i.d.}{\sim} \mathcal{N}(0, 1);$
- ii.  $x_t = -10 (t/n 1/2)^2 + w_t, w_t \stackrel{i.i.d.}{\sim} \mathcal{N}(0, 1);$
- iii.  $x_t = 10 (t/n 1/2)^2 \times w_t, w_t \overset{i.i.d.}{\sim} \mathcal{N}(0, 1);$
- iv.  $x_t = \sum_{i=1}^5 u_i \cos\left(2\pi \left(i/6\right)t\right) + v_i \sin\left(2\pi \left(i/6\right)t\right), u_i, v_i \stackrel{i.i.d.}{\sim} \mathcal{N}\left(0, i/12\right);$
- v.  $x_t = 0.5x_{t-1} + 0.25x_{t-2} + 0.25x_{t-3} + w_t$ ,  $w_t \stackrel{i.i.d.}{\sim} \mathcal{N}(0,1)$ ;

• vi.  $x_t = -x_{t-1} + w_t, w_t \overset{i.i.d.}{\sim} \mathcal{N}(0, 1).$ 

Model	ADF Null without Trend	ADF Null with Trend
i.	Y	Y
ii.	Ν	Y
iii.	Ν	Ν
iv.	Ν	Ν
v.	Y	Y
vi.	Ν	Ν

- i.  $x_t x_{t-1} = w_t w_{t-1}$  is a stationary **ARMA**(p, q) process;
- ii.  $x_t x_{t-1} = 1/n^2 20 (t/n 1/2) / n + w_t w_t$  is not a stationary **ARMA**(p,q) process because of the linear term in t but  $x_t x_{t-1} + bt$  is, where  $b = 20/n^2$ ;
- iii.  $x_t x_{t-1} = 10 \left( t/n 1/2 \right)^2 \times w_t 10 \left( (t-1)/n 1/2 \right)^2 \times w_{t-1}$  has nonconstant variance, so it will not be a stationary **ARMA**(p,q) process.
- iv.  $x_t x_{t-1}$  may be stationary but is not an **ARMA**(p,q) process.
- v.  $x_t x_{t-1} = -0.5 (x_{t-1} x_{t-2}) + 0.25 (x_{t-2} x_{t-3}) + 0.25 (x_{t-3} x_{t-4}) + w_t w_{t-1}$  is a stationary **ARMA**(p, q) process.
- vi.  $x_t x_{t-1} = -(x_{t-1} x_{t-2}) + w_t w_{t-1}$  is a non-stationary **ARMA**(p, q) process.
- (c) Based on (b) and in no more than one sentence, is it possible to detect all kinds of non-stationarity based on an augmented Dickey-Fuller test, with or without trend?

No, some kinds of non-stationarity like non-constant variance will not be detected by a Dickey-Fuller test, regardless of whether or not a trend is included.

## 2. Differencing and Correlation

(a) Express the lag-one autocorrelation function of  $\nabla x_t$  denoted by  $\frac{\gamma_{\nabla x}(1)}{\gamma_{\nabla x}(0)} = \frac{\mathbf{E}[\nabla x_t \nabla x_{t-1}]}{\mathbf{E}[\nabla x_t^2]}$  in terms of the autocovariance function of a mean-zero time series  $x_t$ .

$$\rho_{\nabla x}\left(1\right) = \frac{\gamma_{\nabla x}\left(1\right)}{\gamma_{\nabla x}\left(0\right)} = \frac{\mathrm{E}\left[\nabla x_t \nabla x_{t-1}\right]}{\mathrm{E}\left[\nabla x_t^2\right]} = \frac{2\gamma_x\left(1\right) - \gamma_x\left(0\right) - \gamma_x\left(2\right)}{2\left(\gamma_x\left(0\right) - \gamma_x\left(1\right)\right)}$$

(b) If  $x_t = w_t$ , where  $w_t \stackrel{i.i.d.}{\sim} \mathcal{N}(0,1)$ , then  $\gamma_x(0) = 1$  and  $\gamma_x(h) = 0$  for h > 0. How does the lag-one autocorrelation of  $x_t$  compare to the lag-one autocorrelation of  $\nabla x_t$ ?

$$\rho_{\nabla x}\left(1\right) = -\frac{1}{2}$$

(c) If  $x_t = 0.9x_{t-1} + w_t$ , where  $w_t \stackrel{i.i.d.}{\sim} \mathcal{N}(0, 1 - 0.9^2)$ . Then then  $\gamma_x(h) = 0.9^h$ . How does the lag-one autocorrelation of  $\chi_t$  compare to the lag-one autocorrelation of  $\nabla x_t$ ?

$$\rho_{\nabla x} \left( 1 \right) = \frac{2\left(\frac{9}{10}\right) - 1 - \frac{81}{100}}{2\left(1 - \frac{9}{10}\right)}$$
$$= \frac{\frac{180 - 100 - 81}{100}}{\frac{20}{100}}$$
$$= -\frac{1}{20}.$$

(d) Fill in the blank with "stronger," "weaker," or "possibly stronger or weaker." If a time series  $x_t$  is differenced, the differenced time series  $\nabla x_t$  will have possibly stronger or weaker lag-one autocorrelations than the undifferenced raw time series. This is about how the **magnitudes** of the autocorrelations compare, not the values!

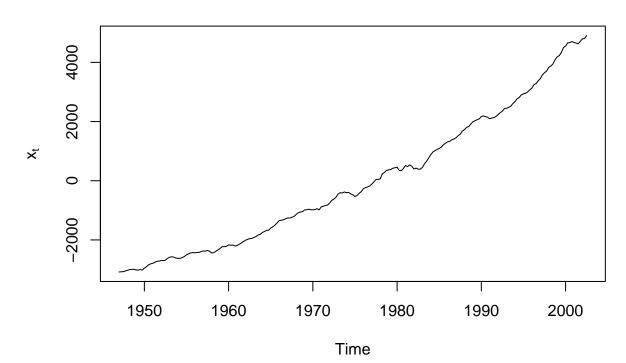
## 3. ARIMA

This question will ask you to analyze the gnp data from the astsa package, which gives the quarterly United States GNP from the first quarter of 1947 to the third quarter of 2002.

```
library(astsa)
data(gnp)
time <- time(gnp)
gnp <- c(gnp) - mean(gnp)
n <- length(gnp)</pre>
```

The data is plotted below.

plot(time, gnp, main = "Quarterly US GNP", ylab = expression(x[t]), xlab = "Time", type = "l")



**Quarterly US GNP** 

(a) Does  $x_t$  appear stationary? If not, do you think  $\nabla x_t$  would appear stationary, based on the plot of the data?

The time series  $x_t$  does not appear stationary because there is a linear trend present, but it looks like  $\nabla x_t$  might be (because differencing will eliminate a linear trend).

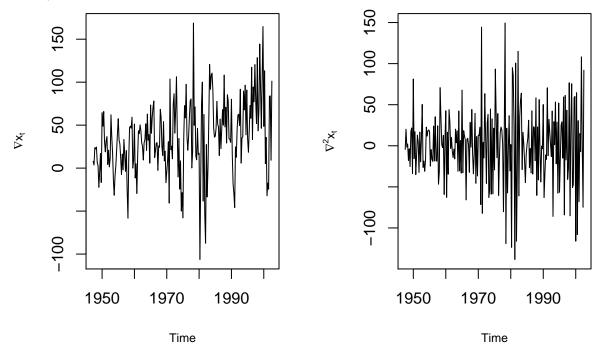
(b) How many times does a level-0.05 augmented Dickey-Fuller test indicate that we should difference the data? The value returned by ndiffs is printed after the code chunk.

```
library(forecast)
ndiffs(gnp, test = "adf", alpha = 0.05, type = "level")
```

## [1] 1

Once.

(c) The first and second differences are plotted below. Based just on the plots alone, how many times do you think we should difference the data?



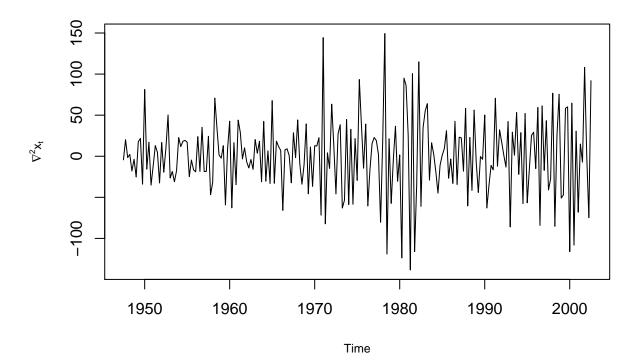
Twice.

(d) Indicate whether or not your conclusions in (b) are the same. If they are not the same, explain how many times **you** think we should difference the data (give a reason for the number of differences you think you should take).

Either once or twice could be accepted as answers, depending on your reasoning!

# 4. ARCH/GARCH

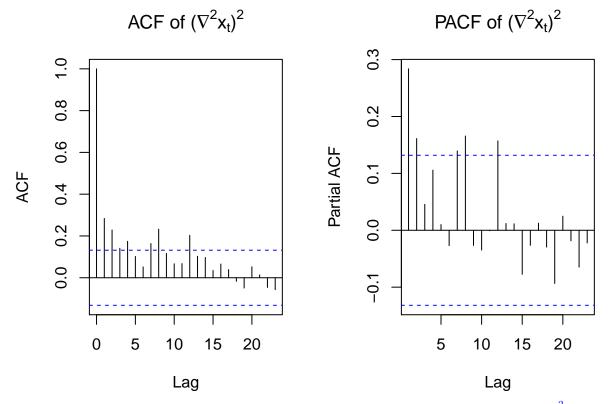
We're going to keep working with the gnp data from the astsa package, and continue to focus on the second differences  $\nabla^2 x_t$ , which are plotted again below.



(a) Based on the above plot, is there evidence of variance nonstationarity?

Yes - the variance appears to be increasing as time passes.

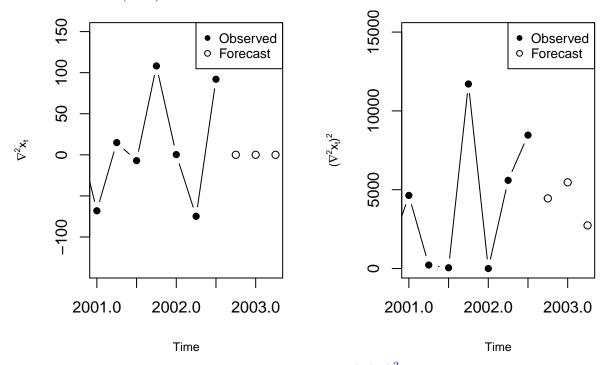
(b) Let's consider a **GARCH**(m, 0) model for  $\nabla^2 x_t$ . Based on the plotted ACF and PACF of the squared second differences  $(\nabla^2 x_t)^2$ , what would you choose for m?



A **GARCH**(m, 0) model for  $\nabla^2 x_t$  corresponds to an **AR**(m) model for the magnitudes  $(\nabla^2 x_t)^2$ . We can select the order of an **AR**(m) model from the PACF by selecting an order equal to the last lag for which the

sample partial autocorrelation is outside of the 95% interval for  $\gamma_x(h) = 0$ . This means we would choose m = 12 here.

(c) Suppose we fit a **GARCH**(m, 0) model for the correct value of m based on part (b) to the second differences. Forecasts of  $\nabla^2 x_t$  and  $(\nabla^2 x_t)^2$  based on this **GARCH**(m, 0) model are given in the plots below. Based the plotted forecasts, does assuming a GARCH model for  $\nabla^2 x_t$  help us forecast  $\nabla^2 x_t$  or the magnitudes  $(\nabla^2 x_t)^2$ ?



Assuming a GARCH model helps us forecast the magnitudes  $(\nabla^2 x_t)^2$ .

(d) In one sentence, explain why your answer in (c) makes sense given what we know about **GARCH**(m, 0) models. Hint: compare the autocorrelation function for  $\nabla^2 x_t$  compared to the autocorrelation function of  $(\nabla^2 x_t)^2$  under a **GARCH**(m, 0) model.

Under a **GARCH**(m, 0) model, the values  $\nabla^2 x_t$  are uncorrelated but the magnitudes  $(\nabla^2 x_t)^2$  are correlated, so past values of  $(\nabla^2 x_t)^2$  help us forecast future values of  $(\nabla^2 x_t)^2$  but past values of  $\nabla^2 x_t$  do not help us forecast future values of  $\nabla^2 x_t$ .

## 5. Spectral Analysis

(a) Suppose

$$\begin{aligned} x_t &= \sum_{k=1}^r v_k \cos\left(2\pi\omega_k t\right) + u_k \sin\left(2\pi\omega_k t\right), \, v_k, u_k \overset{i.i.d.}{\sim} \mathcal{N}\left(0, \sigma_k^2\right) \\ y_t &= \sum_{k=1}^r c_k \cos\left(2\pi\omega_k t\right) + d_k \sin\left(2\pi\omega_k t\right), \, c_k, d_k \overset{i.i.d.}{\sim} \mathcal{N}\left(0, \tau_k^2\right) \end{aligned}$$

The spectral density function of  $x_t$  is  $f(\omega_k) = \sigma_k^2$  and the spectral density function of  $y_t$  is  $g(\omega_k) = \tau_k^2$ .

(a) Write out  $z_t = ax_t + by_t$  in terms of a, b, the  $v_k$ 's, the  $u_k$ 's, the  $c_k$ 's, the  $c_k$ 's, the  $\cos(2\pi\omega_k t)$ 's and the  $\sin(2\pi\omega_k t)$ 's.

$$z_t = ax_t + by_t$$
$$= \sum_{k=1}^r (av_k + bc_k) \cos(2\pi\omega_k t) + (au_k + bd_k) \sin(2\pi\omega_k t)$$

(b) Describe how to define  $e_k$ ,  $f_k$ , and  $\nu_k$  in terms of a, b, the  $v_k$ 's, the  $u_k$ 's, the  $c_k$ 's, and the  $d_k$ 's such that the following holds for  $z_t$ :

$$z_t = \sum_{k=1}^r e_k \cos\left(2\pi\omega_k t\right) + f_k \sin\left(2\pi\omega_k t\right), \ e_k, \ f_k \stackrel{i.i.d.}{\sim} \mathcal{N}\left(0, \nu_k^2\right).$$

$$e_k = av_k + bc_k$$
  

$$f_k = au_k + bd_k$$
  

$$\nu_k = \sqrt{a^2\sigma_k^2 + b^2\tau_k^2}$$

(c) What is the spectral density function  $h(\omega_k)$  of  $z_t$ ?

$$h\left(\omega_k\right) = a^2 \sigma_k^2 + b^2 \tau_k^2.$$

(d) Suppose that

$$\begin{aligned} x_t &= \phi_1 x_{t-1} + w_t, w_t \overset{i.i.d.}{\sim} \mathcal{N}\left(0, \sigma_w^2\right) \\ y_t &= \psi_1 y_{t-1} + v_t, v_t \overset{i.i.d.}{\sim} \mathcal{N}\left(0, \sigma_v^2\right). \end{aligned}$$

Write out  $z_t = ax_t + by_t$ , substituting  $\phi_1 x_{t-1} + w_t$  in for  $x_t$  and  $\psi_1 y_{t-1} + v_t$  in for  $y_t$  and collecting terms that correspond to the same time point.

$$z_t = ax_t + by_t$$
  
=  $a\phi_1 x_{t-1} + b\psi_1 y_{t-1} + aw_t + bv_t.$ 

(e) If we don't assume anything about the values of  $a, b, \phi_1$ , and  $\psi_1$ , can we know whether or not  $z_t$  will be an **AR**(p) process? Hint: If  $z_t$  is an **AR**(p) process, we can find values  $\gamma_1, \ldots, \gamma_p$  such that  $z_t = \gamma_1 z_{t-1} + \cdots + \gamma_p z_{t-p}$ . Just answer yes or no.

No - unless  $\phi_1$  and  $\psi_1$  are equal or a and b are equal, we're not going to be able to write  $z_t$  as a linear function of its past values  $z_{t-1}$ .