

# Notes 5

Maryclare Griffin

3/2/2023

These notes are based on Chapters 1, 5, and 6 of KNNL.

Recall from the previous notes, the linear regression model for a dependent variable or response  $Y$  and independent variables, predictors, or covariates  $X_1, \dots, X_{p-1}$  is defined as:

$$Y_i = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \dots + \beta_{p-1} X_{i,p-1} + \epsilon_i$$

where:

- $\beta_0, \beta_1, \dots, \beta_{p-1}$  are parameters
- $X_{i1}, \dots, X_{i,p-1}$  are known constants
- $\epsilon_i$  is a random error term with mean  $E\{\epsilon_i\} = 0$  and variance  $\sigma^2\{\epsilon_i\} = \sigma^2$ ;  $\epsilon_i$  and  $\epsilon_j$  are uncorrelated so that their covariance is zero (i.e.,  $\sigma\{\epsilon_i, \epsilon_j\} = 0$  for all  $i, j$ ;  $i \neq j$ )
- $i = 1, \dots, n$

Remember, we don't observe  $\beta_0, \beta_1, \dots, \beta_{p-1}$  in the real world. Instead, we **estimate** them by finding the values  $b_0, b_1, \dots, b_{p-1}$  that minimize the sum of squared deviations of the response values  $Y_i$  from the regression function  $\beta_0 + \sum_{k=1}^{p-1} \beta_k X_{ik}$  with respect to  $\beta_0, \beta_1, \dots, \beta_{p-1}$ .

In the previous set of notes, we showed how the `lm` function in `R` can be used to compute  $b_0, b_1, \dots, b_{p-1}$ . We also showed how closed form equations can be derived for  $b_0$  and  $b_1$  when  $p = 2$ . What about when we have multiple predictors? We want to find the values  $b_0, b_1, \dots, b_{p-1}$  that solve all  $p$  of the **normal equations**:

- $\sum_{i=1}^n -2Y_i + 2b_0 + 2\left(\sum_{k=1}^{p-1} b_k X_{ik}\right) = 0$
- For  $k > 0$ ,  $\sum_{i=1}^n -2Y_i X_{ik} + 2b_k X_{ik}^2 + 2b_0 X_{ik} + 2\sum_{l=1, l \neq k}^{p-1} b_l X_{il} X_{ik} = 0$

First, let's rewrite the equations a bit, getting rid of the extra twos and starting with the terms involving elements of  $\mathbf{b}$  and/or  $\mathbf{X}$ :

- $\sum_{i=1}^n b_0 + \sum_{k=1}^{p-1} b_k X_{ik} - Y_i = 0$
- For  $k > 0$ ,  $\sum_{i=1}^n b_k X_{ik}^2 + b_0 X_{ik} + \sum_{l=1, l \neq k}^{p-1} b_l X_{il} X_{ik} - 2Y_i X_{ik} = 0$

For this, we'll need linear algebra. Linear algebra lets us write out sums and sets of equations efficiently. We will define:

- The  $n \times 1$  column vector  $\mathbf{Y} = \begin{pmatrix} Y_1 \\ Y_2 \\ \vdots \\ Y_n \end{pmatrix}$
- The  $n \times p$  **design matrix**  $\mathbf{X} = \begin{pmatrix} 1 & X_{11} & X_{12} & \dots & X_{1,p-1} \\ 1 & X_{21} & X_{22} & \dots & X_{2,p-1} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 1 & X_{n1} & X_{n2} & \dots & X_{n,p-1} \end{pmatrix}$

- The  $p \times 1$  vector  $\mathbf{b} = \begin{pmatrix} b_0 \\ b_1 \\ \vdots \\ b_{p-1} \end{pmatrix}$

**Note:** Matrix multiplication gives us a nice way of computing several sums at once. If  $\mathbf{A}$  is an  $r \times c$  matrix and  $\mathbf{B}$  is a  $c \times s$  matrix, the product  $\mathbf{D} = \mathbf{AB}$  is an  $r \times s$  matrix with elements:

$$d_{ij} = \sum_{k=1}^c a_{ik}b_{kj}.$$

This looks a bit familiar, and suggests that we can write out our normal equations using matrix multiplication. The normal equations involve sums  $\sum_{i=1}^n Y_i X_{ik}$  for different values of  $k$ , as well as sums  $\sum_{i=1}^n b_k X_{ik}^2$ , and  $\sum_{i=1}^n \sum_{l=1, l \neq k}^{p-1} b_l X_{il} X_{ik}$ . The first term resembles what we would expect to obtain if we multiplied  $\mathbf{X}$  and  $\mathbf{Y}$ , however their dimensions are not amenable to matrix multiplication as-is.

**Note:** The transpose of a vector or matrix is obtained by interchanging the rows and columns. \* The transpose of the  $n \times 1$  column vector  $\mathbf{Y}$  is the  $1 \times n$  row vector or matrix  $\mathbf{Y}' = ( Y_1 \ Y_2 \ \dots \ Y_n )$  \* The transpose of the  $n \times p$  matrix  $\mathbf{X}$  is

$$\mathbf{X}' = \begin{pmatrix} 1 & 1 & \dots & 1 \\ X_{11} & X_{21} & \dots & X_{n1} \\ X_{12} & X_{22} & \dots & X_{n2} \\ \vdots & \vdots & \ddots & \vdots \\ X_{1,p-1} & X_{2,p-1} & \dots & X_{n,p-1} \end{pmatrix}$$

We can compute  $\mathbf{X}'\mathbf{Y}$ , let's try it!

$$\mathbf{X}'\mathbf{Y} = \begin{pmatrix} \sum_{i=1}^n Y_k \\ \sum_{i=1}^n Y_k X_{i1} \\ \vdots \\ \sum_{i=1}^n Y_k X_{i,p-1} \end{pmatrix}$$

Aha! We have the  $\sum_{i=1}^n Y_k$  that appears in the first normal equation and the  $\sum_{i=1}^n Y_k X_{ik}$  terms that appear in the remaining  $p - 1$  normal equations. What about the terms involving  $\mathbf{b}$  and elements of  $\mathbf{X}$ ? A natural quantity to consider is  $\mathbf{X}\mathbf{b}$ . Let's compute that!

$$\mathbf{X}\mathbf{b} = \begin{pmatrix} b_0 + \sum_{k=1}^{p-1} b_k X_{1k} \\ b_0 + \sum_{k=1}^{p-1} b_k X_{2k} \\ \vdots \\ b_0 + \sum_{k=1}^{p-1} b_k X_{nk} \end{pmatrix}$$

This isn't quite what we're looking for. We are missing terms involving squares  $X_{ik}^2$  and products  $X_{ik}X_{il}$  for

$l \neq k$ . This suggests seeing what we get if we multiply by  $\mathbf{X}'$  to get a  $p \times 1$  vector:

$$\begin{aligned} \mathbf{X}'\mathbf{X}\mathbf{b} &= \begin{pmatrix} 1 & 1 & \dots & 1 \\ X_{11} & X_{21} & \dots & X_{n1} \\ X_{12} & X_{22} & \dots & X_{n2} \\ \vdots & \vdots & \ddots & \vdots \\ X_{1,p-1} & X_{2,p-1} & \dots & X_{n,p-1} \end{pmatrix} \begin{pmatrix} b_0 + \sum_{k=1}^{p-1} b_k X_{1k} \\ b_0 + \sum_{k=1}^{p-1} b_k X_{2k} \\ \vdots \\ b_0 + \sum_{k=1}^{p-1} b_k X_{nk} \end{pmatrix} \\ &= \begin{pmatrix} \sum_{i=1}^n b_0 + \sum_{k=1}^{p-1} b_k X_{ik} \\ \sum_{i=1}^n X_{i1} \left( b_0 + \sum_{k=1}^{p-1} b_k X_{ik} \right) \\ \sum_{i=1}^n X_{i2} \left( b_0 + \sum_{k=1}^{p-1} b_k X_{ik} \right) \\ \vdots \\ \sum_{i=1}^n X_{i,p-1} \left( b_0 + \sum_{k=1}^{p-1} b_k X_{ik} \right) \end{pmatrix} \\ &= \begin{pmatrix} \sum_{i=1}^n b_0 + \sum_{k=1}^{p-1} b_k X_{ik} \\ \sum_{i=1}^n b_1 X_{i1}^2 + b_0 X_{i1} + \sum_{l=2}^{p-1} b_l X_{i1} X_{il} \\ \sum_{i=1}^n b_2 X_{i2}^2 + b_0 X_{i2} + b_1 X_{i1} X_{i2} + \sum_{l=3}^{p-1} b_l X_{i1} X_{il} \\ \vdots \\ \sum_{i=1}^n b_{p-1} X_{i,p-1}^2 + b_0 X_{i,p-1} + \sum_{l=1}^{p-2} b_l X_{il} X_{i,p-1} \end{pmatrix} \end{aligned}$$

The first element of  $\mathbf{X}'\mathbf{X}\mathbf{b}$  minus the first element of  $\mathbf{X}'\mathbf{Y}$  set equal to 0 gives us the first normal equation! The second element of  $\mathbf{X}'\mathbf{X}\mathbf{b}$  minus the second element of  $\mathbf{X}'\mathbf{Y}$  set equal to 0 gives us the first normal equation! And so on!

**Note:** Two vectors or matrices are said to be equal if they have the same dimension and all of the corresponding elements are equal, i.e. if  $\mathbf{A}$  is an  $r \times c$  matrix and  $\mathbf{B}$  is an  $r \times c$  matrix, then elements of  $\mathbf{A} = \mathbf{B}$  indicates that  $a_{ij} = b_{ij}$  for  $i = 1, \dots, r$  and  $j = 1, \dots, c$ .

**Note:** The difference of two vectors or matrices of the same dimensions is the difference of their elements, i.e. if  $\mathbf{A}$  is an  $r \times c$  matrix and  $\mathbf{B}$  is an  $r \times c$  matrix, then elements of  $\mathbf{D} = \mathbf{A} - \mathbf{B}$  satisfy  $d_{ij} = a_{ij} - b_{ij}$ .

Accordingly, if  $\mathbf{0}$  is a  $p \times 1$  vector with all elements exactly equal to zero, we can write all  $p$  of the normal equations simultaneously using linear algebra as:

$$\mathbf{X}'\mathbf{X}\mathbf{b} - \mathbf{X}'\mathbf{Y} = \mathbf{0}.$$

This equation is a “nice’’ function of  $\mathbf{b}$ ! A few more linear algebra facts will allow us to solve it.

**Note:** The sum of two vectors or matrices of the same dimensions is the sum of their elements, i.e. if  $\mathbf{A}$  is an  $r \times c$  matrix and  $\mathbf{B}$  is an  $r \times c$  matrix, then elements of  $\mathbf{D} = \mathbf{A} + \mathbf{B}$  satisfy  $d_{ij} = a_{ij} + b_{ij}$ .

$$\mathbf{X}'\mathbf{X}\mathbf{b} = \mathbf{X}'\mathbf{Y}$$

**Note:** The **identity matrix** or **unit matrix** is denoted by  $\mathbf{I}$ . It is a diagonal matrix whose elements on the main diagonal  $I_{kk}$  are all equal to 1, and remaining elements are equal to 0. Premultiplying or postmultiplying any  $r \times r$  matrix  $\mathbf{A}$  by the  $r \times r$  identity matrix  $\mathbf{I}$  leaves  $\mathbf{A}$  unchanged, i.e.  $\mathbf{IA} = \mathbf{A}$  and  $\mathbf{A} = \mathbf{AI}$ .

**Note:** The inverse of a  $r \times r$  square matrix  $\mathbf{A}$  is another  $r \times r$  square matrix, denoted by  $\mathbf{A}^{-1}$ , such that  $\mathbf{A}^{-1}\mathbf{A} = \mathbf{I}$ . The inverse matrix  $\mathbf{A}^{-1}$  exists if the matrix  $\mathbf{A}$  is rank  $r$ , i.e. if  $\mathbf{A}$  is **nonsingular** or **full rank**. A matrix with rank less than  $r$  is said to be **singular** or **not of full rank**. The inverse of a matrix with rank  $r$  also has rank  $r$ .

If  $\mathbf{X}'\mathbf{X}$  is full rank, then we have a closed form solution for  $\mathbf{b}$ :

$$\mathbf{b} = (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{Y}$$

**Note:**  $\mathbf{X}'\mathbf{X}$  is full rank when there we cannot write any column of  $\mathbf{X}$  as a linear combination of the remaining columns of  $\mathbf{X}$ . Practically, this means that  $\mathbf{X}'\mathbf{X}$  is never full rank when  $n < p$ .

**Example 1:** Let's return to the data from portrait studios in 21 cities run by Dwaine Studios, Inc. The studios specialize in portraits of children. Let  $X_1$  be the number of persons aged 16 or younger in a city, let  $X_2$  refer to per capita disposable income in a city, and let  $Y$  be the sales of portraits of children in that city from one of the 21 studies. We're going to construct an design matrix and compute  $b_0$ ,  $b_1$ , and  $b_2$  by hand.

```
load("~/Dropbox/Teaching/STAT525/Spring2023/bookdata/dwaine.RData")
X1 <- data$X1 # Extract the first predictor
X2 <- data$X2 # Extract the second predictor
Y <- data$Y # Extract the response
n <- length(Y) # Record the number of observations
X <- cbind(rep(1, n), X1, X2)
XtY <- t(X)%*%Y # Compute X'Y
XtX <- t(X)%*%X # Compute X'X
XtX.inv <- solve(XtX) # Invert XtX
b <- XtX.inv%*%XtY # Solve for b
b0 <- b[1] # Extract the intercept
b1 <- b[2] # Extract b1
b2 <- b[3] # Extract b2
```

We obtain estimates  $b_0 = -68.857$ ,  $b_1 = 1.455$ , and  $b_2 = 9.366$ . These are the same numbers we got from `lm` previously!

Now that we've introduced linear algebra, it's worth noting that we can write the linear regression model and  $Q$ , the sum of squares criterion using linear algebra.

Letting  $\epsilon = \begin{pmatrix} \epsilon_1 \\ \epsilon_2 \\ \vdots \\ \epsilon_n \end{pmatrix}$ , the linear regression model for a dependent variable or response  $Y$  and independent variables, predictors, or covariates  $X_1, \dots, X_{p-1}$  is defined as:

$$\mathbf{Y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}$$

where:

- The elements of  $\boldsymbol{\beta} = \begin{pmatrix} \beta_0 \\ \beta_1 \\ \vdots \\ \beta_{p-1} \end{pmatrix}$  are parameters
- The elements of the  $n \times p$  matrix  $\mathbf{X} = \begin{pmatrix} 1 & X_{11} & X_{12} & \dots & X_{1,p-1} \\ 1 & X_{21} & X_{22} & \dots & X_{2,p-1} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 1 & X_{n1} & X_{n2} & \dots & X_{n,p-1} \end{pmatrix}$  are known constants

- For  $i = 1, \dots, n$ ,  $\epsilon_i$  is a random error term with mean  $E\{\epsilon_i\} = 0$  and variance  $\sigma^2\{\epsilon_i\} = \sigma^2$ ;  $\epsilon_i$  and  $\epsilon_j$  are uncorrelated so that their covariance is zero (i.e.,  $\sigma\{\epsilon_i, \epsilon_j\} = 0$  for all  $i, j; i \neq j$ )

The  $p \times 1$  vector of estimated regression coefficients  $\mathbf{b}$  corresponds to the value of  $\beta$  that minimizes

$$Q = (\mathbf{Y} - \mathbf{X}\beta)'(\mathbf{Y} - \mathbf{X}\beta)$$

Accordingly, the  $p$  normal equations are obtained by taking the derivative of  $Q$  with respect to  $\beta$  and setting them equal to  $\mathbf{0}$ . We can see this by first expanding out  $Q$ , and then taking derivatives.

$$Q = \mathbf{Y}'\mathbf{Y} - \mathbf{Y}'\mathbf{X}\beta - (\mathbf{X}\beta)'\mathbf{Y} + (\mathbf{X}\beta)'(\mathbf{X}\beta)$$

**Note:** Given two  $1 \times p$  vectors  $\mathbf{a}$  and  $\mathbf{c}$ ,  $\mathbf{a}'\mathbf{c} = \mathbf{c}'\mathbf{a}$ .

$$Q = \mathbf{Y}'\mathbf{Y} - 2\mathbf{Y}'\mathbf{X}\beta + (\mathbf{X}\beta)'(\mathbf{X}\beta)$$

**Note:** The transpose of a product of matrices  $(\mathbf{AB})'$  is obtained by reversing the order of the elements in the product and taking the transpose of each,  $\mathbf{B}'\mathbf{A}'$ .

$$Q = \mathbf{Y}'\mathbf{Y} - 2\mathbf{Y}'\mathbf{X}\beta + \beta'\mathbf{X}'\mathbf{X}\beta$$

**Note:** Given an  $1 \times p$  vector  $\mathbf{A}$ , the derivative of  $\mathbf{A}\beta$  with respect to  $\beta$  is  $\mathbf{A}$ .

$$\frac{\partial Q}{\partial \beta} = -2\mathbf{Y}'\mathbf{X} + \frac{\partial Q}{\partial \beta}\beta'\mathbf{X}'\mathbf{X}\beta$$

**Note:** Given an  $1 \times p$  vector  $\mathbf{A}$ , the derivative of  $\beta'\mathbf{A}\beta$  with respect to  $\beta$  is  $2\beta'\mathbf{A}$ .

$$\frac{\partial Q}{\partial \beta} = -2\mathbf{Y}'\mathbf{X} + 2\beta'\mathbf{X}'\mathbf{X}$$

If we substitute  $\mathbf{b}$  in for  $\beta$ , this looks very similar to the expression we got for the normal equations but transposed and multiplied by 2. Substituting  $\mathbf{b}$  in for  $\beta$  and taking the transpose, we get:

$$-2\mathbf{X}'\mathbf{Y} + 2\mathbf{X}'\mathbf{X}\mathbf{b} = \mathbf{0}$$

Dividing by 2 and rearranging yields our normal equations:

$$\mathbf{X}'\mathbf{X}\mathbf{b} - \mathbf{X}'\mathbf{Y} = \mathbf{0}$$

Ta-da!